

Numerical Optimization for Physicists and Statisticians

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Grateful to many physicist collaborators:

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N. Michel, W. Nazarewicz, F.M. Nunes, E. Olsen, T. Papenbrock,
P.-G. Reinhardt, N. Schunck, M. Stoitsov, J. Vary, K. Wendt, and others

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Mathematical/Numerical Optimization for ISNET

Possible Topics Today

- Optimization Basics
- Optimization for Expensive Model Calibration
 - fast, limiting the number of expensive simulation evaluations
 - local, given enough resources, find you a point for which you cannot improve the objective in a local neighborhood

derivative-free - useful in situations where derivatives unavailable

- $^{\diamond}$ Beyond χ^2 Minimization
- Stochastic Optimization
- Bayesian Optimization

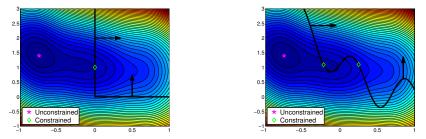
1. Mathematical/Numerical Nonlinear Optimization

Optimization is the "science of better"

Find parameters (controls) $x = (x_1, \ldots, x_n)$ in domain Ω to improve objective f

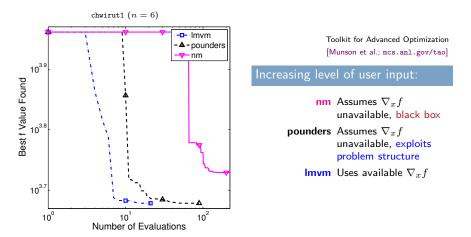
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\min\left\{f(x):x\in\Omega\subseteq\mathbb{R}^n\right\}
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- $^{\diamond}$ (Unless Ω is very special) Need to evaluate f at many x to find a good \hat{x}_{*}
- ♦ Focus on local solutions: $f(\hat{x}_*) \leq f(x) \ \forall x \in \mathcal{N}(\hat{x}_*) \cap \Omega$



Implicitly assume that uncertainty modeled through constraints and objective(s)

The Price of Algorithm Choice: Solvers in PETSc/TAO

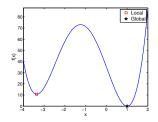


Observe: Constrained by budget on #evals, method limits solution accuracy/problem size

Why Not Global Optimization, $\min_{x \in \Omega} f(x)$?

Careful:

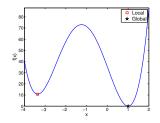
- $^{\diamond}$ Global convergence: Convergence (to a local solution/stationary point) from anywhere in Ω
- $^{\diamond}$ Convergence to a global minimizer: Obtain x_{*} with $f(x_{*}) \leq f(x) \, \forall x \in \Omega$



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Anyone selling you global solutions when derivatives are unavailable:

either assumes more about your problem (e.g., convex f)

or expects you to wait forever Törn and Žilinskas: An algorithm converges to the global minimum for any continuous f if and only if the sequence of points visited by the algorithm is dense in Ω .

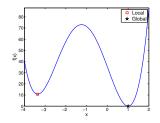
or cannot be trusted

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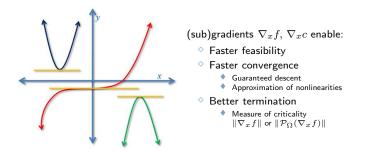
Instead:

- $^{\diamond}$ Rapidly find good local solutions and/or be robust to poor solutions
- Find several good local solutions concurrently (APOSMM/LibEnsemble)

Optimization Tightly Coupled With Derivatives (WRT Parameters)

Typical optimality (no noise, smooth functions)

$$\nabla_x f(x_*) + \lambda^T \nabla_x c_E(x_*) = 0, c_E(x_*) = 0$$



But derivatives $\nabla_x S(x)$ are not always available/do not always exist

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Obtain Derivatives $\nabla_x S$ Whenever Possible

Handcoding (HC)

- "Army of students/programmers"
 - ? Prone to errors/conditioning
 - ? Intractable as number of ops increases

Algorithmic/Automatic Differentiation (AD)

"Exact* derivatives!"

- ? No black boxes allowed
- ? Not always automatic/cheap/well-conditioned

Finite Differences (FD)

"Nonintrusive"

- ? Expense grows with n
- ? Sensitive to stepsize choice/noise

 \rightarrow [Moré & W.; SISC 2011], [Moré & W.; TOMS 2012]

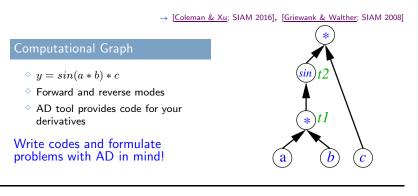
... then apply derivative-based method (that handles inexact derivatives)







Algorithmic Differentiation



Many tools (see www.autodiff.org):

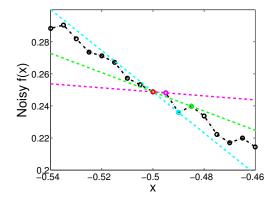
F OpenAD F/C Tapenade, Rapsodia C/C++ ADOL-C, ADIC Matlab ADiMat, INTLAB Python/R ADOL-C

Also done in AMPL, GAMS, JULIA!

Numerical Differentiation

The Problem: Finite differences sensitive to choice of h

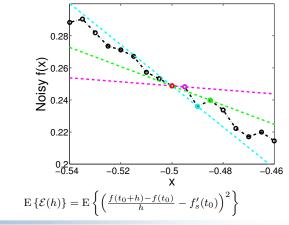
$$\frac{f(t_0+h) - f(t_0)}{h} \approx f'_s(t_0)$$



Numerical Differentiation

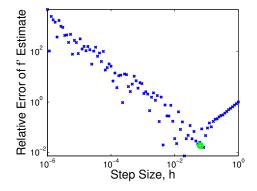
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Minimize

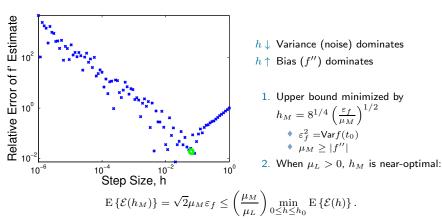
Optimal Forward Difference Parameter h



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Optimal Forward Difference Parameter h



 $\frac{1}{4}\mu_L^2 h^2 + 2\frac{\varepsilon_f^2}{h^2} \le \mathbf{E}\left\{\mathcal{E}(h)\right\} \le \frac{1}{4}\mu_M^2 h^2 + 2\frac{\varepsilon_f^2}{h^2}$

[Estimating Noisy Derivatives. Moré & W., TOMS 2012]]

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Simulation-Based Optimization

$$\min_{x \in \mathbb{R}^n} \left\{ f(x) = F[\mathbf{S}(\mathbf{x})] : c(\mathbf{S}(\mathbf{x})) \le 0, x \in \mathcal{B} \right\}$$

Optimize expensive, nonlinear functions arising in science & engineering

"parameter estimation", "model calibration", "design optimization", \ldots

- $^{\diamond}~f:\mathbb{R}^n\to\mathbb{R}$ objective, $S:\mathbb{R}^n\to\mathbb{R}^p$ numerical simulation, Ω constraints
- $^{\diamond}\,$ Evaluating S means running a simulation modeling some (smooth) process
 - Ex- S = solving PDEs via finite elements
 - \blacklozenge Here: assume f is from a deterministic computer simulation
- $^{\diamond}~S$ can contribute to objective and/or constraints, possibly noisy
- $^{\diamond}\,$ Derivatives $\nabla_x S$ often unavailable or prohibitively expensive to obtain/approximate directly
- $^{\diamond}~S$ (could/must be parallelized) takes secs/mins/hrs/days for 1 x

Evaluation is a bottleneck for optimization

B compact, known region (e.g., finite bound constraints)



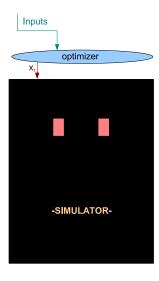
Deterministic Algorithms

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"Simplest" (=Most Naive) Formulation: Blackbox f



Optimizer gives x, physicist provides f(x)

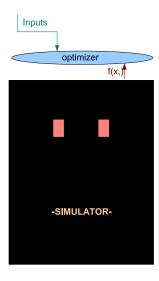
- f can be a blackbox (executable only or proprietary/legacy codes)
- Only give a single output
 - no derivatives with respect to x: $\nabla_x S(x), \nabla^2_{x,x} S(x)$
 - no problem structure

Good solutions guaranteed in the limit, but:

Computational budget limits number of evaluations

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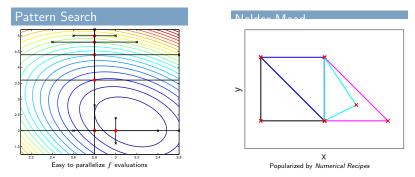
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Two main styles of local algorithms

- Direct search methods (pattern search, Nelder-Mead, ...)
- Model- ("surrogate-")based methods (quadratics, radial basis functions, ...)

Algorithms: Direct Search Methods



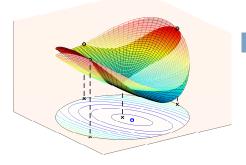
 \diamond Rely on indicator functions: $[f(x_k + \mathbf{s}) < f(x_k)]$

 \diamond Work with **black-box** f(x), do not exploit structure F[x, S(x)]

→ [Kolda, Lewis, Torczon, SIREV 2003]

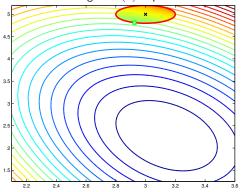
Trust-Region Methods Use Models Instead of f

To reduce the number of expensive f evaluations \rightarrow Replace difficult optimization problem $\min f(x)$ with a much simpler one $\min \{m(x) : x \in B\}$



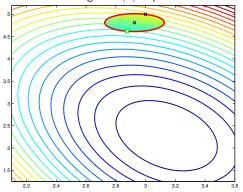
Classic NLP Technique:

- *f* Original function: computationally expensive, no derivatives
- m Surrogate model: computationally attractive, analytic derivatives



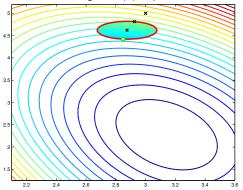
Use a surrogate m(x) in place of the unwieldy f(x)

- [◊] Trust *m* to approximate *f* within $\mathcal{B} = \{x \in \mathbb{R}^n : ||x - x_k|| \le \Delta_k\},\$
- ◇ Obtain next point from $\min \{m(x) : x \in B\}$,
- Evaluate function and update (x_k, Δ_k) based on how good the model's prediction was.



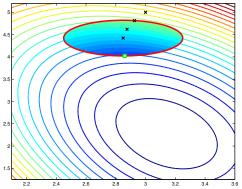
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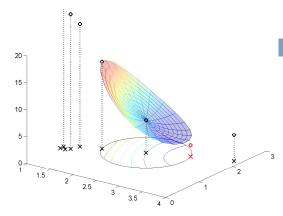
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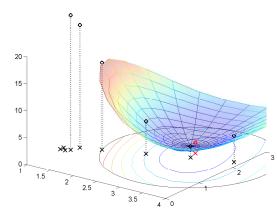


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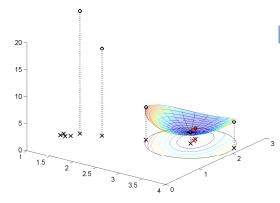
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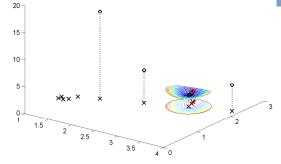
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- $^{\diamond}$ Trust m_k within region \mathcal{B}_k
- $^{\diamond}\,$ Minimize m_k within \mathcal{B}_k to obtain next point for evaluation
- Do expensive evaluation
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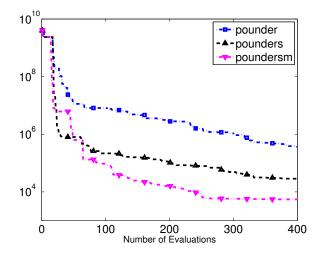
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Exploit Structure!

Performance of Model-Based Methods



Optimizing EDF in [Bertolli et al., PRC 2012]

Parameter Estimation is NOT a Blackbox Problem

Generic:

$$\min_{x} \left\{ f(x) : x \in \Omega \subseteq \mathbb{R}^n \right\}$$

- \boldsymbol{x} n decision variables
- $f~:\mathbb{R}^n \rightarrow \mathbb{R}$ objective function
- Ω feasible region,
 - $\{x: c_E(x) = 0, c_I(x) \le 0\}$
 - c_E (vector of) equality constraints
 - c_I (vector of) inequality constraints

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Typical calibration problem:

$$f(x) = \|\mathbf{R}(x)\|_2^2 = \sum_{i=1}^p R_i(x)^2$$

- x n coupling constants
- $\begin{array}{l} R_i \ : \mathbb{R}^n \to \mathbb{R} \ \text{residual function} \\ \mathsf{Ex.-} \ \ \frac{1}{w_i} \left(S(x; \theta_i) d_i \right) \\ & \blacklozenge \ \ S(x; \theta_i): \ \text{numerical simulation} \\ \mathsf{Ex.-} \ \ \mathsf{Obtain} \ \chi^2(x) \ \text{by} \ \ \frac{1}{p-n} f(x) \end{array}$

$$\Omega = \{ x : \mathbf{l} \le x \le \mathbf{u} \}$$

- Finite bounds (for some x_i)
- Often dictated by dom(S)

[Ekström et al, PRL 2013] [Kortelainen et al, PRC 2014]

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 $^{\diamond}$ Taking advantage of structure should further reduce # of expensive evaluations

Exploiting Nonlinear Least Squares Structure

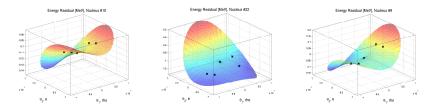
Obtain a vector of output $R_1(x), \ldots, R_p(x)$

 \diamond (Locally) Model each R_i by a surrogate $q_k^{(i)}$

$$R_i(x) \approx q_k^{(i)}(x) = R_i(x_k) + (x - x_k)^\top \mathbf{g}_k^{(i)} + \frac{1}{2}(x - x_k)^\top \mathbf{H}_k^{(i)}(x - x_k)$$

Employ models in the approximation

$$\begin{array}{lll} \nabla f(x) & = \sum_i \nabla \mathbf{R}_i(\mathbf{x}) R_i(x) & \longrightarrow \sum_i g_k^{(i)}(x) R_i(x) \\ \nabla^2 f(x) & = \sum_i \nabla \mathbf{R}_i(\mathbf{x}) \nabla \mathbf{R}_i(\mathbf{x})^T + R_i(x) \nabla^2 \mathbf{R}_i(\mathbf{x}) & \longrightarrow \sum_i g_k^{(i)}(x) g_k^{(i)}(x)^T + R_i(x) \mathbf{H}_k^{(i)}(x) \end{array}$$



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General Nonlinear Least Squares

$$\min_{x} f(x) = \|\mathbf{R}(x)\|_{\mathbf{W}}^2$$

$$\begin{split} \mathbf{R} &: \mathbb{R}^n \to \mathbb{R}^p \text{ "residual vector"} \\ &\to \mathsf{Think:} \ R_i(x) = S(x; \theta_i) - d_i \\ \mathbf{W} \text{ norm: } \|\mathbf{y}\|_{\mathbf{W}} = \left(\mathbf{y}^T \mathbf{W} \mathbf{y}\right)^{1/2} \\ &\to \mathbf{W} = I_p \text{ recovers } \|\cdot\|_2 \end{split}$$

General Nonlinear Least Squares

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General Nonlinear Least Squares

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$$\begin{split} \mathbf{W} \quad \text{symmetric positive definite} \\ &\bullet \mathbf{W} = \mathbf{W}^T \\ &\bullet \mathbf{y}^T \mathbf{W} \mathbf{y} > 0 \text{ for all } \mathbf{y} \neq \mathbf{0} \end{aligned}$$

$$f(x) = \sum_{i=1}^p \sum_{j=1}^p W_{i,j} R_i(x) R_j(x) \ge 0 \end{split}$$

• $\mathbf{W} = (\operatorname{diag}(\sigma))^{-1}$ yields familiar

$$f(x) = \sum_{i=1}^{p} \frac{(S(x;\theta_i) - d_i)^2}{\sigma_i} = \sum_{i=1}^{p} \frac{R_i(x)^2}{\sigma_i}$$



A Warning

Can I pass this to my favorite $\min_x \chi^2(x) = \|\tilde{\mathbf{R}}(x)\|^2$ solver?

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$$\sum_{i=1}^{p} \sum_{j=1}^{p} \left(\tilde{R}_{i,j}(x) \right)^{2}$$

= $\sum_{i=1}^{p} \sum_{j=1}^{p} \left(\sqrt{|W_{i,j}R_{i}(x)R_{j}(x)|} \right)^{2}$
\ne $\sum_{i=1}^{p} \sum_{j=1}^{p} W_{i,j}R_{i}(x)R_{j}(x)$

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Can I pass this to my favorite $\min_x \chi^2(x) = \|\tilde{\mathbf{R}}(x)\|^2$ solver?

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= $\sum_{i=1}^{p} \sum_{j=1}^{p} \left(\sqrt{|W_{i,j}R_{i}(x)R_{j}(x)|} \right)^{2}$
\neq $\sum_{i=1}^{p} \sum_{j=1}^{p} W_{i,j}R_{i}(x)R_{j}(x)$

- ! Allow for complex-valued residuals
- ! Disallow $W_{i,j}R_i(x)R_j(x) < 0$

In any case, you will likely suffer algorithmically

Relationship to Covariance Matrices

Data $\{(\theta_1, d_1), \cdots, (\theta_p, d_p)\}$

 $^{\diamond}$ Errors independent and normally distributed: $d \sim N(\mu, \Sigma)$,

 $d_i = \mu(\theta_i; x_*) + \varepsilon_i, \qquad \varepsilon_i \sim N(0, \sigma_i^2) \qquad i = 1, \dots, p.$

 Σ is a $p \times p$ diagonal matrix, with *i*th diagonal entry σ_i^2

 \diamond Model, $S(\theta; x)$ with Gaussian errors:

$$\left[S(\theta_1; x), \cdots, S(\theta_p; x)\right]^T \sim N\left(\mu(\cdot; x), C\right),$$

 $\circ C$ a $(p \times p$ symmetric positive definite) covariance matrix accounting for correlation between model outputs (i.e., $Cov(S(\theta_i; x), S(\theta_j; x)) = C_{i,j})$

Assuming model errors are independent of data errors,

$$[m(\hat{x};\theta_1) - d_1, \cdots, m(\hat{x};\theta_p) - d_p]^T \sim N(0, C + \Sigma),$$

$$\diamond$$
 Joint likelihood $l(x; \theta; d) \propto \exp\left[-\frac{1}{2}\mathbf{R}(x; \theta)^T \left(\mathbf{C} + \mathbf{\Sigma}\right)^{-1} \mathbf{R}(x; \theta)\right]$

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 Σ is a $p \times p$ diagonal matrix, with *i*th diagonal entry σ_i^2

 \diamond Model, $S(\theta; x)$ with Gaussian errors:

$$\left[S(\theta_1; x), \cdots, S(\theta_p; x)\right]^T \sim N\left(\mu(\cdot; x), C\right),$$

 $\circ C$ a $(p \times p$ symmetric positive definite) covariance matrix accounting for correlation between model outputs (i.e., $Cov(S(\theta_i; x), S(\theta_j; x)) = C_{i,j})$

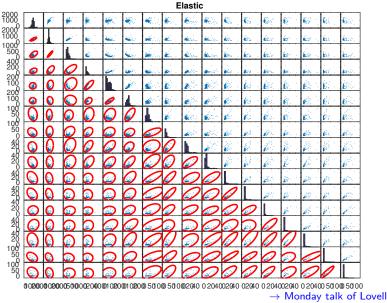
Assuming model errors are independent of data errors,

$$[m(\hat{x};\theta_1) - d_1, \cdots, m(\hat{x};\theta_p) - d_p]^T \sim N(0, C + \Sigma),$$

$$^{\circ}$$
 Joint likelihood $l(x; \theta; d) \propto \exp\left[-\frac{1}{2}\mathbf{R}(x; \theta)^T \left(\mathbf{C} + \mathbf{\Sigma}\right)^{-1} \mathbf{R}(x; \theta)\right]$

Warning: C, Σ can no longer hide behind constants of proportionality

Optical Potentials: Incorporating Covariances in \boldsymbol{W}



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Applications Using the Jacobian $[\hat{J}]_{i,j} = \frac{\partial R_i(\hat{x})}{\partial x_j} = \frac{1}{w_i} \frac{\partial S(x;\theta_i)}{\partial x_j}$

Residual $\mathbf{R}(x) \in \mathbb{R}^p$ undergoes a change by $\boldsymbol{\epsilon} \in \mathbb{R}^p$

$$^{\diamond}$$
 Ex.- normalized datum $rac{d_i}{w_i}$ is changed to $rac{d_i}{w_i}+\epsilon_i$

 $\hat{\mathbf{x}} \in \arg\min_{\hat{\mathbf{x}} \in \mathbb{R}^n} f^0(x) = \|\mathbf{R}(x)\|_2^2 \qquad \hat{\mathbf{x}}_{\epsilon} \in \arg\min_{\hat{\mathbf{x}} \in \mathbb{R}^n} f(x) = \|\mathbf{R}(x) + \epsilon\|_2^2$

A second-order expansion of $f = \|\mathbf{R}(x) + \boldsymbol{\epsilon}\|_2^2$ about $\hat{\mathbf{x}}$:

$$f(\hat{\mathbf{x}}) + 2\epsilon^T \hat{J}(x - \hat{\mathbf{x}}) + \frac{1}{2}(x - \hat{\mathbf{x}})^T \left(\nabla^2 f^0(\hat{\mathbf{x}}) + 2\sum_{i=1}^p \epsilon_i \nabla^2 R_i(\hat{\mathbf{x}})\right)(x - \hat{\mathbf{x}}),$$

When ϵ is small, this quadratic will be convex and hence minimized at

$$x_{\epsilon} - \hat{\mathbf{x}} = 2 \left(\nabla^2 f^0(\hat{\mathbf{x}}) \right)^{-1} \hat{J}^T \boldsymbol{\epsilon} + \mathcal{O}(\|\boldsymbol{\epsilon}\|^2).$$

When $\mathbf{R}(\hat{\mathbf{x}})$ is small, $\nabla^2 f^0(\hat{\mathbf{x}})\approx 2\hat{J}^T\hat{J}$ and

$$\tilde{x}_{\epsilon} \approx \hat{\mathbf{x}} + \left(\hat{J}^T \hat{J}\right)^{-1} \hat{J}^T \boldsymbol{\epsilon}$$

Stochastic Optimiza

Or

Stochastic Optimization

General problem

$$\min\left\{f(x) = \mathbb{E}_{\xi}\left[F(x,\xi)\right]: \ x \in X\right\}$$
(1)

- $^{\diamond} \ x \in \mathbb{R}^n$ decision variables
- $^{\diamond}$ ξ vector of random variables
 - $\bullet \ \text{independent of } x$
 - $P(\xi)$ distribution function for ξ
 - ξ has support Ξ
- $^{\diamond}\ F(x,\cdot)$ functional form of uncertainty for decision x
- $^{\diamond}~X\subseteq \mathbb{R}^{n}$ set defined by deterministic constraints

Approach of Sampling Methods for $f(x) = \mathbb{E}_{\xi} \left[F(x, \xi) \right]$

- $\diamond \ \, {\rm Let} \ \xi^1,\xi^2,\cdots,\xi^N\sim P$
- \diamond For $x \in X$, define:

$$f_N(x) = \frac{1}{N} \sum_{i=1}^N F(x, \xi^i)$$

f_N is a random variable (really, a stochastic process)

(depends on
$$\left(\xi^1,\xi^2,\cdots,\xi^N\right)$$
)

• Motivated by $\mathbb{E}_{\xi} [f_N(x)] = f(x)$

Bias of Sampling Methods

$$^{\diamond}\;\; {\rm Let}\; f^* = f(x^*) \; {\rm for}\; x^* \in X^* \subseteq X$$

Bias of Sampling Methods

$$^{\diamond}$$
 Let $f^* = f(x^*)$ for $x^* \in X^* \subseteq X$

 \diamond For any $N \ge 1$:

$$\mathbb{E}_{\xi}\left[f_{N}^{*}\right] \leq f^{*} = \mathbb{E}_{\xi}\left[F(x^{*},\xi)\right]$$

because

$$\mathbb{E}_{\xi}\left[f_{1}^{*}\right] = \mathbb{E}_{\xi}\left[\min\left\{F(x,\xi): x \in X\right\}\right] \le \min\left\{\mathbb{E}_{\xi}\left[F(x,\xi)\right]: x \in X\right\} = f^{*}$$

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- $^{\diamond}$ Sampling problems result in optimal values below f^{*}
- $^{\diamond}~f_{N}^{*}$ is biased estimator of f^{*}

Sample Average Approximation

- \diamond Draw realizations $\hat{\xi}^1, \hat{\xi}^2, \cdots, \hat{\xi}^N \sim P$ of $\left(\xi^1, \xi^2, \cdots, \xi^N\right)$
- \diamond Replace (1) with

$$\min\left\{\frac{1}{N}\sum_{i=1}^{N}F(x,\hat{\xi}^{i}):\ x\in X\right\}$$
(2)

- $\hat{f}_N(x) = \frac{1}{N} \sum_{i=1}^N F(x, \hat{\xi}^i)$ deterministic
- Follows mean of the N sample paths defined by the (fixed) $\hat{\xi}^i$

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Convergence with ${\boldsymbol N}$

A sufficient condition:

For any ε > 0 there exists N_ε so that

$$\left|\hat{f}_N(x) - f(x)\right| < \epsilon \qquad \forall N \ge N_\epsilon \quad \forall x \in X$$

with probability 1 (wp1).

- $^{\diamond}$ Then $\hat{f}_N^* \to f^*$ wp1.
- ◇ (With additional assumptions on f and $X^* \subset X$):

 ${\rm dist}(x_N^*,X^*)\to 0$

◇ (+ uniqueness, $X^* = x^*$):

 $x_N^* \to x^*$

Stochastic Approximation Method

Basically just:

Input x^0 1. $x^{k+1} \leftarrow \mathcal{P}_X \{x^k - \alpha_k s^k\},$ $k = 0, 1, \dots$

- $\diamond \alpha_k$ a step size
- $^{\diamond} s^k$ a random direction

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Generally assume:

$$\begin{split} &\alpha_{k} \colon \sum_{k=0}^{\infty} \alpha_{k} = \infty, \sum_{k=0}^{\infty} \alpha_{k}^{2} < \infty \qquad (\text{e.g., } \alpha_{k} = \frac{c}{k}) \\ &s^{k} \colon \mathrm{E}\left\{\nabla f(x^{k})^{T} s^{k}\right\} > 0 \\ &s^{k} \text{ is an ascent direction (in expectation) at } x^{k} \end{split}$$

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 $\begin{aligned} \alpha_k \colon \sum_{k=0}^{\infty} \alpha_k &= \infty, \sum_{k=0}^{\infty} \alpha_k^2 < \infty \\ s^k \colon E\left\{\nabla f(x^k)^T s^k\right\} > 0 \\ s^k \text{ is an ascent direction (in expectation) at } x^k \end{aligned}$ (e.g., $\alpha_k = \frac{c}{k}$)

$$\diamond$$
 "Exact" Stochastic Gradient Descent: $s^k = \nabla f(x^k)$

Classic SA Algorithms

- "Original" method is Robbins-Monro (1951)
- Without derivatives: Kiefer-Wolfowitz (1952) replaces gradient with finite-difference approximation, e.g.,

1.
$$x^{k+1} \leftarrow x^k - \alpha_k s^k$$
, $k = 0, 1, \dots$

where

$$s^{k} = \frac{F(x^{k} + h_{k}I_{n}; \hat{\xi}^{k}) - F(x^{k} - h_{k}I_{n}; \hat{\xi}^{k+1/2})}{2h_{k}}$$

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- Requires 2n evaluations every iteration
- Can appeal to variance reduction techniques (e.g., common RNs)
- Convergence $x^k \to x^*$ if f strongly convex (near x^*), usual conditions on α_k ,

$$h_k \to 0$$
, $\sum_k \frac{\alpha_k^2}{h_k^2} < \infty$

• K-W recommend:
$$\alpha_k = \frac{1}{k}$$
, $h_k = \frac{1}{k^{1/3}}$

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$$\alpha_k = \frac{1}{k}$$
, $h_k = \frac{1}{k^{1/3}}$

 Extensions such as SPSA (Spall) reduce number of evaluations (see randomized methods slides...)

Derivative-Based Stochastic Gradient Descent

Input x^0 ; Repeat:

- 1. Draw realization $\hat{\xi}^k \sim P$ of ξ^k 2. Compute $s^k = \nabla_x F(x^k; \hat{\xi}^k)$ 3. Update $x^{k+1} \leftarrow \mathcal{P}_X \{x^k \alpha_k s^k\}$

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- $\hat{\nabla}_x F(x^k; \hat{\xi}^k)$ is an unbiased estimator for $\nabla f(x^k)$

Derivative-Based Stochastic Gradient Descent

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- $\nabla_x F(x^k; \hat{\xi}^k)$ is an unbiased estimator for $\nabla f(x^k)$
- Can incorporate curvature if desired
 - e.g., $B^k s^k$ an unbiased estimator for $\left(\nabla^2 f(x^k)\right)^{-1} \nabla f(x^k)$
- Can work with subgradients
- \diamond Can even output $x^N = \frac{1}{N} \sum_{k=1}^N x^k$

Randomized Algorithms for Deterministic Problems

 $\min\left\{f(x): x \in X \subseteq \mathbb{R}^n\right\}$

- \diamond f deterministic
- Random variables are now generated by the method, not from the problem
- Often assume properties of f
 e.g., ∇f is L'-Lipschitz:

$$\|\nabla f(x) - \nabla f(y)\| \le L' \|x - y\| \qquad \forall x, y \in X$$

e.g., f is strongly convex (with parameter τ):

$$f(x) \ge f(y) + (x - y)^T \nabla f(y) + \frac{\tau}{2} ||x - y||^2 \quad \forall x, y \in X$$

Basic Algorithms

Matyas (e.g., 1965):

> Input x⁰; repeat:
1. Generate Gaussian u^k (centered about 0)
2. Evaluate
$$f(x^k + u^k)$$
3. $x^{k+1} = \begin{cases} x^k + u^k & \text{if } f(x^k + u^k) < f(x^k) \\ x^k & \text{otherwise.} \end{cases}$

Basic Algorithms

Matyas (e.g., 1965):

> Input x⁰; repeat:
1. Generate Gaussian u^k (centered about 0)
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3. x^{k+1} =
$$\begin{cases} xk + uk & \text{if } f(xk + uk) < f(xk) \\ xk & \text{otherwise.} \end{cases}$$

Poljak (e.g., 1987)

◇ Input x⁰, {h_k, µ_k}_k; repeat:
1. Generate a random u^k ∈ Rⁿ
2. x^{k+1} = x^k - h_k
$$\frac{f(x^k + \mu_k u^k) - f(x^k)}{\mu_k} u^k$$

• h_k > 0 is the step size

• $\mu_k > 0$ is called the smoothing parameter

Applying SA-Like Ideas to Special Cases

$$\min\left\{f(x) = \frac{1}{m}\sum_{i=1}^{m}F_i(x) : x \in X\right\}$$

 $m \ \mathsf{huge}$

Applying SA-Like Ideas to Special Cases

$$\min\left\{f(x) = \frac{1}{m}\sum_{i=1}^{m}F_i(x) : x \in X\right\}$$

m huge

- Ex.- Nonlinear Least Squares $F_i(x) = \|\phi(x; \theta^i) - d^i\|^2$ Evaluating $\phi(\cdot, \cdot)$ requires solving a large PDE
- Ex.- Sample Average Approximation $F_i(x) = R(x; \hat{\xi}^i)$ $\hat{\xi}^i \in \Omega$ a scenario/RV realization (and R depends nontrivially on $\hat{\xi}^i$)

Warning: likely nonconvex!

Applying SA-Like Ideas to Special Cases

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- $\begin{array}{l} \text{Ex.-} \quad \textit{Nonlinear Least Squares} \\ F_i(x) = \|\phi(x;\theta^i) d^i\|^2 \\ \text{Evaluating } \phi(\cdot,\cdot) \text{ requires solving a large PDE} \end{array}$
- Ex.- Sample Average Approximation $F_i(x) = R(x; \hat{\xi}^i)$ $\hat{\xi}^i \in \Omega$ a scenario/RV realization (and R depends nontrivially on $\hat{\xi}^i$)

The good:

 $\nabla f(x) = \sum_{i=1}^{m} \nabla F_i(x)$

The bad:

 $\diamond m$ still huge

Warning: likely nonconvex!

$$\min\left\{f(x) = \frac{1}{m}\sum_{i=1}^{m}F_i(x): x \in X\right\}$$

" $F_i(x)$ is a member of a population of size m"

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 $^{\diamond}$ Randomly sample \mathcal{S} , a subset of size $|\mathcal{S}|$, from $\{1,\cdots,m\}$

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 $^\diamond~~ {\rm Use}~ -\nabla f_{\mathcal{S}} = -\frac{1}{|\mathcal{S}|}\sum_{i\in\mathcal{S}} \nabla F_i(x)$ as direction s^k

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$$\diamond~$$
 Use $-\nabla f_{\mathcal{S}}=-\frac{1}{|\mathcal{S}|}\sum_{i\in\mathcal{S}}\nabla F_i(x)$ as direction s^k

 $^{\diamond}$ How to choose \mathcal{S} ?

$$\mathbb{E}\left\{\left\|\nabla f_{\mathcal{S}_n} - \nabla f\right\|^2\right\} = \left(1 - \frac{|\mathcal{S}|}{m}\right) \mathbb{E}\left\{\left\|\nabla f_{\mathcal{S}_r} - \nabla f\right\|^2\right\}$$

 \Rightarrow sampling without replacement (S_n) gives lower variance than does sampling with replacement (S_r)

Bayesian Optimization for Approximate Global Optimization

Statistical approaches (e.g., EGO [Jones et al., 1998])

- enjoy global exploration properties,
- excel when simulation is expensive, noisy, nonconvex
- ... but offer limited support for constraints

[Schonlau et al., 1998]; [Gramacy & Lee, 2011]; [Williams et al., 2010]

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Combine (global) statistical (objective-only) optimization tools

a) response surface modeling/emulation: training a flexible model f^k on $\{x^{(i)},y^{(i)}\}_{i=1}^k$ to guide choosing $x^{(k+1)}$

e.g., [Mockus, et al., 1978], [Booker et al., 1999]

- b) expected improvement (EI) via Gaussian process (GP) emulation [Jones, et al., 1998]
- ... with a tool from mathematical programming
 - c) augmented Lagrangian (AL): for handling nonlinear constraints [Powell, 1969], [Bertsekas, 1982], ...

Similar approach for combining other data terms

[Picheny, Gramacy, W., Le Digabel. NIPS 2016]; [Gramacy et al, Technometrics 2016]

Expected Improvement

$$\label{eq:intermediate} \text{Improvement:} \qquad I(x) = \max\{0, f_{\min}^k - Y(x)\}, \qquad f_{\min}^k \equiv \min_{i=1,\dots,k} f(x^i)$$

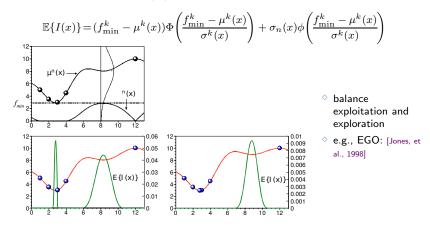
Expectation of improvement (EI) has closed-form expression:

$$\mathbb{E}\{I(x)\} = (f_{\min}^k - \mu^k(x))\Phi\left(\frac{f_{\min}^k - \mu^k(x)}{\sigma^k(x)}\right) + \sigma_n(x)\phi\left(\frac{f_{\min}^k - \mu^k(x)}{\sigma^k(x)}\right)$$

Expected Improvement

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Separate, Independent Component Modeling

$$\circ f \longrightarrow Y_f(x)$$

 $\circ c = (c_1, \dots, c_m) \longrightarrow Y_c(x) = (Y_{c_1}(x), \dots, Y_{c_m}(x))$

Distribution of composite random variable serves as a surrogate for $L_A(x; \lambda, \rho)$:

$$Y(x) = Y_f(x) + \lambda^{\top} Y_c(x) + \frac{1}{2\rho} \sum_{j=1}^{m} \max(0, Y_{c_j}(x))^2$$

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Simplifications when f is known:

Composite posterior mean available in closed form; e.g., under GP priors:

$$\mathbb{E}\{Y(x)\} = \mu_f^k(x) + \lambda^\top \mu_c^k(x) + \frac{1}{2\rho} \sum_{j=1}^m \mathbb{E}\{\max(0, Y_{c_j}(x))^2\}$$

Generalized EI [Schonlau et al., 1998] gives

$$\mathbb{E}\{\max(0, Y_{c_j}(x))^2\} = \sigma_{c_j}^{2n}(x) \left[\left(1 + \left(\frac{\mu_{c_j}^k(x)}{\sigma_{c_j}^k(x)} \right)^2 \right) \Phi\left(\frac{\mu_{c_j}^k(x)}{\sigma_{c_j}^k(x)} \right) + \frac{\mu_{c_j}^k(x)}{\sigma_{c_j}^k(x)} \phi\left(\frac{\mu_{c_j}^k(x)}{\sigma_{c_j}^k(x)} \right) \right]$$

Summary

- Move beyond "blackbox" optimization
- Exploiting structure vields better solutions, in fewer simulations
- Promote optimization/modeling considerations during code development
- Correlated residuals a first step
- Highlights attention that must be paid to model and data uncertainties
- Can repeat for nonGaussian, MAPs,

[www.mcs.anl.gov/tao (Optimization toolkit) www.mcs.anl.gov/~wild (Get in touch!)]

Grateful to relevant coauthors

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Thank You!